

Today

1) Probability Review

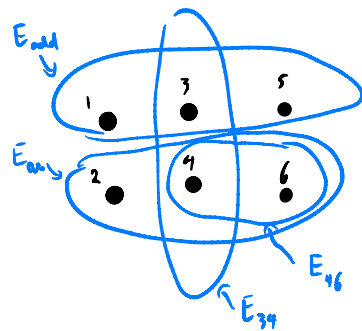
2) Fun w/ Gaussians

(Discrete) Probability Review

A discrete probability space consists of

- 1) A ^(countable) set of "outcomes" Ω called the sample space
- 2) A probability function $P: \Omega \rightarrow [0, 1]$

$$\sum_{\omega \in \Omega} P(\omega) = 1$$



An event E is a subset of Ω

The probability of E is $P(E) := \sum_{\omega \in E} P(\omega)$

$$P(\omega) = 1/6 \quad \forall \omega \in \Omega$$

$$E_1 \cup E_2 \leftrightarrow \text{"}E_1 \text{ or } E_2\text{"}$$

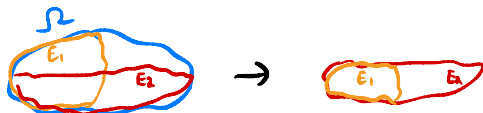
$$E_1 \cap E_2 \leftrightarrow \text{"}E_1 \text{ and } E_2\text{"}$$

The probability of E_1 conditional on E_2 is $P(E_1 | E_2) := \frac{P(E_1 \cap E_2)}{P(E_2)}$



$$\text{E.g. } P(E_{46} | E_{\text{even}}) = \frac{2}{3} \quad \text{and} \quad P(E_{345} | E_{\text{odd}}) = \frac{1}{3}$$

Events E_1 and E_2 are independent if $P(E_1 \cap E_2) = P(E_1) \cdot P(E_2)$



E_{34}, E_{even} independent

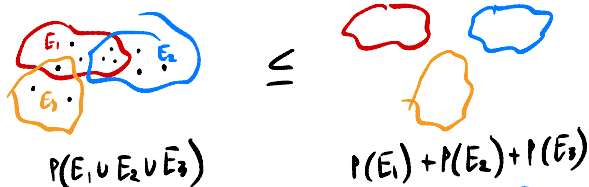
$$P(E_1 | E_2) \cdot P(E_2) = P(E_1) \cdot P(E_2)$$

$$\downarrow$$

$$P(E_1 | E_2) = P(E_1)$$

The Union Bound

Thm: Given ^{"bad"} events E_1, E_2, \dots, E_k $P(E_1 \cup E_2 \cup \dots \cup E_k) \leq \sum_{i=1}^k P(E_i)$



Proof: Let $n_0 := |\{i : E_i \cap \emptyset \neq \emptyset\}|$

$$\text{Thm. } P(E_1 \cup E_2 \cup \dots) = \sum_{\emptyset \in E_1 \cup E_2 \cup \dots \cup E_k} P(\emptyset)$$

$$\leq \sum_{\emptyset \in E_1 \cup E_2 \cup \dots \cup E_k} P(\emptyset) \cdot n_0$$

$$= \sum_{\emptyset} \sum_{i: \emptyset \in E_i} P(\emptyset)$$

$$= \sum_{i=1}^k \sum_{\emptyset \in E_i} P(\emptyset) \quad (\text{Switched order of summation})$$

$$= \sum_{i=1}^k P(E_i)$$

Random Variables

A discrete RV is a function $X: \Omega \rightarrow \mathbb{R}$

" $X=a$ " is event $\{\omega \in \Omega: X(\omega)=a\}$

Random variables X, Y are independent if $\Pr(X=a \cap Y=b) = \Pr(X=a) \cdot \Pr(Y=b) \forall a, b \in \mathbb{R}$

$X := \text{dice \#}$
 $Y := \mathbb{1} \text{ (dice is odd)}$
 $Z := X + 10(1-Y)$

1	1	3	3	5	5
•	1	•	1	•	1
2	12	4	14	6	16
•	0	•	0	•	0

Expectation of RV X is $\mathbb{E}[X] := \sum_i i \cdot \Pr(X=i)$ \rightarrow E.g. $\mathbb{E}[Y] = \frac{51}{6}$

Conditional Expectation of RV X conditioned on $Y=y$ is

$$\mathbb{E}[X|Y=y] := \sum_i i \cdot \Pr(X=i|Y=y)$$

Variance of RV X is $\text{Var}(X) := \mathbb{E}[(X - \mathbb{E}[X])^2]$ ($\sigma(X) = \sqrt{\text{Var}(X)}$ is standard dev.)

Covariance of RVs X, Y is $\text{Cov}(X, Y) := \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])]$

Common Identities (Simple Calculations)

$$E[aX + bY] = aE[X] + bE[Y]$$

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

$$\text{Var}(cX) = c^2 \cdot \text{Var}(X)$$

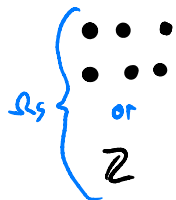
If X, Y independent

$$E[XY] = E[X] \cdot E[Y]$$

$$\text{Cov}(X, Y) = 0$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

Discrete RV



$$\Pr(X=a)$$

$$\mathbb{E}[X] = \sum_i i \cdot \Pr(X=i)$$

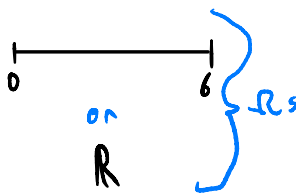
$$\Pr(X \leq a) = \sum_{i \leq a} \Pr(X=i)$$

$$\sum_i \Pr(X=i) = 1$$

$$\text{Events} = 2^{\mathbb{R}}$$

Probability
Function

Continuous RV



$$\text{pdf } f(a)$$

$$\mathbb{E}[X] := \int_i i \cdot f(i) di$$

$$\text{CDF } \Phi(a) := \int_{-\infty}^a f(i) di$$

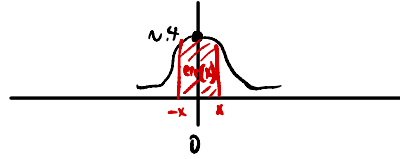
$$\int_{-\infty}^{\infty} f(i) di = 1$$

$$\sigma\text{-algebra} \subseteq 2^{\mathbb{R}}$$

Measure

Gaussians

$$\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$



$Z \sim N(0,1) \rightarrow$ "standard Gaussian"

$E[X] = 0$ by symmetry around y-axis

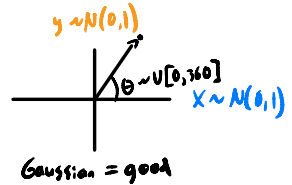
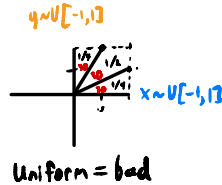
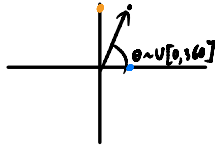
Amazing Fact 1: $\int_{-\infty}^{\infty} \phi(x) = 1$ (no anti-deriv. w/ ele. fns.)

$$\text{erf}(x) = \frac{1}{\sqrt{\pi}} \int_{-x}^x e^{-t^2}$$

Can derive $\text{var}(X) = 1$ from this (hence "N(0,1)")

Amazing Fact 2: Rotational symmetry

Suppose want uniformly random direction in \mathbb{R}^2 w/ independent coordinates



Fix (a,b) w/ length $\sqrt{a^2+b^2}$

$$\Pr(V=(a,b)) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(a^2+b^2)}$$

The same probability $\forall u=(x,y)$ s.t. $\sqrt{x^2+y^2} = \sqrt{a^2+b^2}$



Generalizes to \mathbb{R}^k for $k > 2$

Defn. $N(\mu, \sigma^2) = \mu + \sigma Z$ where $Z \sim N(0, 1)$ is a "non-standard Gaussian"

If $X \sim N(\mu, \sigma^2)$ then $\mathbb{E}[X] = \mu$ and $\text{var}(X) = \sigma^2 \rightarrow \text{Prove}$

Amazing Fact 3: Sum of Gaussians is a Gaussian

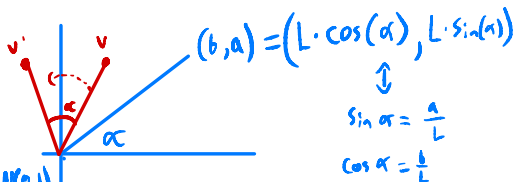
If $X \sim N(0, 1)$, $Y \sim N(0, 1)$, X, Y independent, $a, b \in \mathbb{R}$
then $aX + bY \sim N(0, a^2 + b^2)$

Let $L = \sqrt{a^2 + b^2}$

Consider random vector V'

1) $V \sim (X, Y)$ where $X, Y \sim N(0, 1)$

2) $V' \leftarrow V$ rotated α degrees



V' has same distribution as V (by rot. symm.)

V' has coordinates $(\sim, \frac{X \cdot \sin \alpha + Y \cdot \cos \alpha}{L})$ (by trig.)
 $\sim N(0, 1)$

$$\rightarrow \frac{a}{L} \cdot X + \frac{b}{L} \cdot Y \sim N(0, 1)$$

$$\rightarrow aX + bY \sim L \cdot N(0, 1) = N(0, a^2 + b^2)$$

Amazing Fact 4: Central Limit Theorem

Suppose X_1, X_2, \dots, X_n are i.i.d. w/ $\mathbb{E}[X_i] = \mu$, $\text{var}(X_i) = \sigma^2$

Let $S_n = \sum_1^n X_i$ (so $\mathbb{E}[S_n] = n\mu$ and $\text{var}(S_n) = n \cdot \sigma^2$)

Let $\hat{S}_n = \frac{S_n - n\mu}{\sqrt{n} \cdot \sigma}$ (so $\mathbb{E}[\hat{S}_n] = 0$ and $\text{var}(\hat{S}_n) = 1$)

Then $\hat{S}_n \rightarrow Y$ where $Y \sim N(0, 1)$

$$\Downarrow$$
$$\Pr(\hat{S}_n \leq a) = \Pr(Y \leq a) \pm o_n(1) \quad \forall a \in \mathbb{R}$$